



TI 2001-036/3
Tinbergen Institute Discussion Paper

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Version: April 2002

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Dynamics of China's Regional Development and Pollution

An investigation into the Environmental Kuznets Curve

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Abstract

This paper addresses the existence of an Environmental Kuznets Curve for China, using a sample of thirty regions, covering the period 1982–1997. The types of pollution included are wastewater, waste gas and solid waste. We consider the development of the sources of pollution in a pooled cross-section analysis with pollution in absolute levels, in per capita terms and relative to real Gross Regional Product (GRP). At intermediate levels of GRP per capita, the increase of solid and gas emissions tends to decelerate, but accelerates again at high levels of GRP per capita. Water pollution decreases with per capita GRP. We also predict future waste gas emissions.

Key-words: Environmental Kuznets Curve, China, pollution, economic development

JEL-codes: O1, Q0

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1. Introduction

China's economy has developed rapidly over the past 20 years. At the same time serious environmental problems have emerged. In this paper, we focus on the relationship between China's regional development and regional pollution. Following the seminal work by Grossman and Krueger (1995)² we address the question whether an Environmental Kuznets Curve (EKC) – i.e., a hump shaped relationship between per capita income and pollution – can be found for China. Simultaneously studying the development of per capita income and pollution in China is interesting for at least two reasons. The first is related to the size of the Chinese economy and its population. The fast economic development and the size of the economy will result in tremendous environmental impacts of the Chinese economy on both local and global environmental problems in the very future (see also Sachs, 1997). The second reason relates to the specific difficulties that arise in many studies based on cross-country analysis. These studies face the problem that the cross-sectional units can be relatively difficult to compare (or to be made comparable). Income data for high-income countries are more accurate than those for low-income countries (see Stern et al., 1996). By focusing on Chinese data, the quality of the data can be expected to yield no bias in the estimation results (without arguing that data are reliable per se). Also, the problems arising from using different – potentially inconsistent – databases with a different sample of countries do not occur in the case of China (see, for example, Grossman and Krueger, 1995, as compared to Shafik and Bandyopadhyay, 1992). Furthermore, for cross-country studies per capita income data have to be denominated in a common currency. This is done by either ‘deflating’ using market exchange rates or purchasing power parities. It is known that these different methods yield diverging results, in particular when very poor or former communist countries are involved with high import barriers (e.g., Temple, 1999). By focusing on China, this problem does not arise, although it has to be recognised that purchasing power differs over the regions. Finally, China has a relatively uniform political system, which reduces the problem of unobserved heterogeneity.

For these reasons, the present study can make a useful contribution to the research on the existence of an Environmental Kuznets Curve. The availability of data for 30 regions over the period 1982–1997 allows for a pooled cross-section analysis in which the position of the Environmental Kuznets Curve can be studied, allowing for fixed effects differing across regions. Such a study can be relevant even in a relatively homogeneous country such as China as patterns of specialisation among regions can be of major influence due to, e.g., participation in world trade

² Without aiming at providing an exhaustive list of references in the field, some of the relevant early studies that followed up on Grossman and Krueger (1995) are Panayotou (1993), Selden and Song (1994), Shafik and Bandyopadhyay (1992), and Stern et al. (1996).

or the effects of foreign direct investments. In addition the present study can contribute to a better prediction of the future development of Chinese environmental problems under different development scenarios.

The main conclusion of the paper is that it is virtually impossible to talk about *the* Kuznets curve. It turns out that the relationship between pollution and income is highly dependent on the type of pollution that is considered and on how environmental impact is being measured (that is, in terms of levels of pollution, pollution per capita or pollution per unit of real gross regional product (GRP)). For example, for solid waste measured in absolute levels we find an N-shaped relationship, while there is no statistically significant relationship if we consider emissions per capita, and it is downward sloping if pollution is modelled per unit of gross national product. Furthermore, statistically significant differences in region-specific fixed effects are found implying that the location of the Kuznets curve is not stable across regions. This implies substantial region-specific components in pollution that may be related to, for example, the specific industrial structure resulting from regional specialisation patterns.

The outline of this paper is as follows. In Section 2, we describe China's regional development. The environmental problems and the distribution of pollution among regions are discussed in Section 3. In Section 4, we discuss the main mechanisms that determine the shape of the Income-Emission Relationship. We continue with an analysis of the empirically observed relationship between regional development and regional pollution in Section 5. Section 6 concludes.

2. China's regional development

In this section we describe the uneven regional development in China during the past two decades in order to provide a background relevant for the interpretation of the results from the econometric analysis. Unlike most developing countries, China's economy has grown rapidly. From 1986 to 1997, the average annual growth rate of gross domestic product (GDP) was 11.2%.³ The average standard of living has increased considerably as well. However, there exist major differences between regions.

2.1 Regional development disparity

To illustrate regional disparities in China we use two provinces as an example. In 1978, Guangdong's GRP per capita was 367 yuan. In the same year Gansu's GRP per capita was 348

³ Unless explicitly stated otherwise, the data used in this paper are from China Statistical Yearbooks.

yuan. In 1997, Guangdong's GRP per capita amounted to 10428 yuan, whereas Gansu's GRP per capita had risen to only 3137 yuan. A more systematic overview is provided by making a distinction between coastal (in the sequel denoted by *) and landlocked regions and between northern (N) and southern (S) regions. Figures A1, A2 and A3 in Appendix 1 provide details about the classification of the regions. Table A1 in the appendix provides regional growth rates for different time periods. For the period 1978–1993 the average (unweighted) growth rate of GRP in northern regions was 9.1% and it was 10.1% for southern regions. For the coastal and landlocked regions the annual growth rates were 10.4% and 9.0%, respectively. During the period 1994–1997 the northern and the southern regions experienced annual growth rates of 10.9% and 12.8%, respectively. For the coastal and landlocked regions these growth rates were 12.9% and 11.2%, respectively. Additional figures about the disparity between coastal and landlocked areas and its evolution over time are presented in Table 1 below.

< Insert Table 1 about here >

Further data are provided in Table 2. They reveal that the North is – on average – relatively sparsely populated, relatively rich in terms of GRP per capita in 1978, but falling behind over time, specialising towards agriculture and away from industrial output.

< Insert Table 2 about here >

2.2 Causes of the disparity of regional growth rates

Most Chinese economists agree on the factors that have contributed to the increase of the regional economic disparity since 1978 (see Hu et al., 1995, Wei Houkai, 1997, and Zhou Minliang, 1998).

First, the regional disparity is mainly attributable to different paces of development of non state-owned enterprises (Zhou Minliang, 2000). In the northern part of China, non state-owned enterprises contributed 64.1% to the gross output value of industry, while in the southern part this was 76.4%. In 1996, 72.6% of total foreign investments were allocated to the southern part of China. Guangdong, Jiangsu and Shanghai attracted 48.9% of total FDI. In the same year, 83.2% of the contribution by Hong Kong, Macao and Taiwan to the gross industrial output value originated in the south.

Second, the changes in industrial structure differ by region. At the outset of the policy of economic reform, heavy industry was mainly located in northern China, whereas southern China

mainly had light industry. In 1981, the gross output value of light industry in northern China was 45.0% of total industrial output value, in southern China it was 57.1%. In the beginning of the 1980s, when a shortage of light industry products became manifest, the central government initiated preferential policy to encourage the light industry. The enterprises in southern China followed the market and benefited. In 1996, refrigerators, electric fans, household washing machines, radios, recorders, TV sets, and cameras produced in Southern China amounted to 50.8%, 95.5%, 57.9%, 98.0%, 97.7%, 79.6% and 84.0% of total China's output, respectively. The price reform of heavy industry products came relatively late, thereby hindering the development of northern China.

Third, the central government's policy has favored coastal areas, especially those in southern China. In 1978 China started to reform and opened its economy to the world. During the period 1979–1983, the central government provided more financial means to two southern coastal provinces, Guangdong and Fujian, in order to support the market-oriented reform. These provinces were also entitled to absorb investments from abroad. Subsequently four special economic zones were established, 14 coastal cities were opened to trade, three delta regions were declared open economic regions, and Hainan became a special economic province. In 1988, the strategy of coastal area development was proposed. The main idea was to develop an export-oriented economy in coastal areas.

In addition to these specific regional policies, the central government dismantled the organisationally inefficient collective farming. It encouraged foreign investments and allowed for the development of non state-owned enterprises, adjusted the price system and descended some economic decision powers to local governments. These policies produced strong incentives to the development of all provinces, but especially of those in southern coastal areas.

3. China's environmental problems in a regional perspective

China's regional development is characterised by some socially undesirable features. First, there are large income inequalities among regions, which have widened over time. Second, serious unemployment prevails in areas where state-owned heavy industries are located. Third, regions suffer from social problems: especially in periods of price increases, people in less developed regions are more heavily affected in their living standards than those living in advanced regions. Finally, the combination of these (negative) tendencies has resulted in accumulated environmental problems, often concentrated in specific areas. In the remainder of this section, we elaborate on these environmental problems. Other reviews on China's pollution problems can be found in Vermeer (1998) and Edmonds (1999).

3.1 Soil erosion and water problems

Soil erosion occurs in 38% of the country's area. About 2.62 million square km is desertified. It is estimated that as a consequence 15–20% of all species of plants and animals is endangered. Cultivated land reduces by 100 thousand ha annually. Many scientists argue that the overexploitation of forests was an important cause of the flood appearing in the Chang River in 1998. Today China's forested area per capita is only 1/9 of the average world level. In Hainan, the natural forest reduced from 12 million ha in the 1960s to 5.3 million ha at present. Water shortage is becoming serious, especially in northern China, due to little rainfall. About 300 cities in China face water shortage, which is deemed urgent in 50 cities. In Gansu, though the government has built up a forested region for water resources preservation, forests are still in decline and the snow line is rising because of inefficient protection and drought (Li Xihui, 1999).

3.2. Pollution

Table 3 and Figures A4–A6 provide insight into the discharge levels of industrial wastewater, industrial waste gas emissions (defined as emissions of CO₂, NO₂ and SO₂) and industrial solid waste disposal. This evidence clearly displays the regional variation in industrial pollution.

< Insert Table 3 about here >

In view of the fact that the coastal areas comprise only 13.4% of the country's surface, we see that the pollution problem is heavily concentrated there. It is also clear that industrial pollution is increasing rapidly, except for wastewater. Hence, China's environmental situation regarding pollution is deteriorating (see, for example, also the National Environment Protection Agency, 1996). The following casual observations underline this judgement. In 1995, water quality in the Yangtze River was – to Chinese standards – very poor over more than 24% of its length. For the Yellow River, Peril River, Huai River, Hai River, Songhua and Liao River these figures are 60%, 22%, 51%, 41% and 67%, respectively. More than 70% of the surface water (rivers) near cities is polluted. More than 50% of the groundwater in the cities is polluted to different extents. Xinhuashe (1999) states that “At present, there are 31 rivers and 51 sewing systems dumping sewage into the Bohai Sea. Every year, about 2.8 billion tons of sewage is discharged, accounting for 33% of the country's total. The discharged solid waste reaches 700,000 tons, 50% of the coastal areas' total in this category, while the area of the Bohai Sea is only one sixtieth of the State's total sea area”. Of all 600 cities, less than 1% have acceptable air quality. Recently, the

World Resources Institute listed the 10 cities in the world with the highest air pollution; nine of them are in China. Acid rain problems occur in 30% of all areas. It is estimated that the direct loss by air and water pollution amounts to 4–8% of China's GDP annually (Xinhuashe, 1999). In the remainder of this section, we discuss the regional situation concerning wastewater, waste gas and solid waste in more detail.

3.2.1 Industrial wastewater

In 1997, 18.8 billion tons of industrial wastewater was discharged. Industrial wastewater was mainly produced (63.8%) by four sectors: petroleum processing and cokes production, smelting and pressing of ferrous metals, raw chemical materials and chemical products, and the paper industry. From 1991 to 1997, the total volume of industrial wastewater emissions has declined in most provinces, but there are several provinces or autonomous regions where emissions of industrial wastewater have increased.

3.2.2 Industrial waste gas

Chinese statistical sources only report total emissions of industrial waste gas and of SO₂. Separate data on CO₂ and NO₂ are not available. Industrial waste gas mainly originates from three sectors: production and supply of electric power, gas and water, smelting and pressing of non-ferrous metals, and non-metal mineral products. They account for 72% of China's total volume of industrial waste gas emission in 1997. From 1991 to 1997, the country's industrial waste gas emissions increased by 34%. Industrial waste gas emissions consist of two parts: waste gas originating from the process of fuel burning and waste gas from other production processes. In 1997, the former part was responsible for 63% of total industrial waste gas emissions. Waste gas from production and supply of electric power, gas and water accounted for 58% of this. Of the latter part, waste gas from non-metal mineral production, especially cement production and smelting and pressing of non-ferrous metals, constituted almost 70% of total waste gas in the production process.

Waste gas shows the same regional trend as industrial wastewater in that it is highly concentrated. Six provinces (Hebei, Liaoning, Jiangsu, Shandong, Henan and Sichuan) are responsible for 41% of total emissions. There is more particulate matter in the cities of northern China than in the South, which – combined with waste gas, and given an arid or semi-arid environment – increases air pollution. In southern areas, there is a gradual extension of areas suffering from acid rain due to the presence of more low hills, abundant rainfall and wet climate, mixed with increasing waste gas emissions.

Other emissions also reflect the regional disparity. In 1997, China emitted 13.63 million tons of SO₂. Eight provinces (Hebei, Shanxi, Liaoning, Jiangsu, Shandong, Henan, Sichuan and Guizhou) were responsible for 54% of the total and Shandong for 11%. Of the country's 6.85 million tons of industrial soot emissions in 1997, Hebei, Liaoning, Heilongjiang, Shandong, Henan and Sichuan were responsible for 41%. Of the total 5.48 million tons of industrial dust emissions, six provinces (Hebei, Liaoning, Shandong, Henan, Guangdong, and Sichuan) caused 41%.

3.2.3 Industrial solid waste

Of all industrial solid waste, 86% originated from three industrial sectors: mining and quarrying, production and supply of electric power, water and gas and, finally, smelting and pressing of non-ferrous metals. The regional disparity in industrial solid waste is different from the pattern for industrial wastewater and industrial waste gas. Table 3 shows that about 57% of solid waste is produced in the landlocked regions. Eight provinces (Hebei, Shanxi, Liaoning, Jiangxi, Shandong, Sichuan, Heilongjiang and Henan) caused 55% of the country's industrial solid waste emissions. The growth rates are highest in Beijing, Neimenggu, Shaanxi and Xinjiang.

3.3 Pollution damage

Pollution damage is related to many factors, such as the coverage by plants, annual rainfall, the share of cultivated land, and the usage of technologies. Plants can absorb waste gas, rainfall can dilute wastewater, and advanced technology can improve pollution abatement. In China, there are relatively many areas consisting of plateaux, mountains and deserts. Only a small percentage of land is cultivated and population is concentrated in a few regions. The little coverage of forest (less than 15%) can explain the low absorption of waste gas. The limited seasonal rainfall, especially in North China, enhances stock pollution. The relatively poor technology makes it difficult to reduce pollution effectively. In addition, pollution is concentrated in densely populated areas.

4. The relationship between regional development and regional pollution

In recent years, the Environmental Kuznets Curve (EKC) hypothesis has received much attention in the environmental economics literature. The literature builds on the seminal article by Kuznets (1955) in which he derived a hump-shaped relationship between per capita income and income inequality. In 1965 – after Kuznets had published his article to depict the relationship between income distribution and economic development – Williamson (1965) published a paper focusing

on regional inequality and economic development. He analysed data from 24 countries, and found that whereas the regional gap in low-income countries is lower, the regional gap in middle income countries is higher, and that in high-income countries it is lower again. Thus he obtained the (hump shaped) Kuznets curve and concluded that as the economy develops, the regional gap will experience an increase first and a decline afterwards.

The hypothesis of the Environmental Kuznets Curve claims that as an economy develops, environmental problems will initially get more serious, but will eventually decline. The hypothesis postulates a relationship between welfare and environmental quality and might lead to a better understanding of the opportunities for sustainable development. For the analysis that is to follow, it is important to emphasise that the inverted U-shaped relationship between per capita income and pollution is but one of the many forms that a more general Income-Emission Relationship can take. An inverted U-curve need not suffice to describe the relationship between environmental and economic development. The executive director of the UN Population Fund once said: “Much of the environmental degradation witnessed today is due primarily to two groups of people – the top billion richest and the bottom billion poorest” (Todaro, 1997). This suggests that for high incomes, the relationship between pollution and income is positive again. Or, even stronger, that the environmental Kuznets curve is U-shaped. A motivation for the latter can be as follows. Environmental deterioration can be split into two parts: pollution and the reduction of natural resources (land, forest, grass and mineral resources). The U-curve reflects the fact that environmental degradation in underdeveloped countries is related to population pressure, extensive production modes, and the overexploitation of natural resources. In advanced countries it is more related to excessive consumption. The export of many developing countries is mainly concentrated in raw materials and imports are mainly industrial products, while advanced countries import and consume raw materials. Though the top richest and bottom poorest people are both the main origins of environmental degradation, the developing countries often bear the worst environmental consequences. Given this wide range of reasonable possibilities, we will in our empirical application allow for a flexible specification of the income-emission relationship that allows for non-linearities between pollutants and per capita income.

To understand the development of environmental quality in China, it is relevant to keep in mind three factors that affect the relationship between per capita income and emissions. The *scale* effect entails that in a growing economy emissions tend to increase (at given emission intensities and a given industrial structure). Given the high growth rates that China experienced in the past decades, this has an important positive effect on emissions. The *intensity* effect is affected by several factors, all of which are related to technology (see, e.g., Garbaccio et al, 1999, for

empirical information on the development of the Chinese energy-output ratio). For that reason it is sometimes also called technique effect (see Copeland and Taylor, 2002). As economies grow rich, they can afford more advanced and more efficient technologies resulting in lower intensities. This may be one mechanism that gives rise to a ‘real’ negative relationship between income and emissions. However, intensities are also affected by changing prices of (or taxes on) polluting inputs into the production process. These changes in price are unlikely to be directly related to per capita income, but they may have important impacts on emissions. Arguments have been put forward that, as people become wealthier, they tend to push for stricter environmental policies (for example, de Bruyn and Heintz, 1999). This may give rise to a relationship between per capita income and emissions along the effects of per capita income on, for example, taxes on pollution or environmental standards.⁴ Still another factor that may affect the pollution intensities of technologies are foreign direct investments. This may be particularly relevant for certain regions in China. Unfortunately, however, no accurate data are available to test for this. Third, one has to consider the *structural* (or *composition*) effect. A typical time path of the sectoral composition of an economy is one in which countries are initially characterised by a large agricultural sector, followed by a period of industrialisation and subsequently followed by de-industrialisation and a rising service sector (for example, Baumol, 1967, Maddison, 1991, and De Groot 2000). It is beyond the scope of this paper to discuss the underlying mechanisms in detail. However, one can conclude from the literature on sectoral developments that as consumers get richer, they initially tend to shift their consumption pattern towards manufacturing products and at later stages towards services. Under the assumption that services are pollution-extensive and manufacturing goods pollution-intensive, such a demand-driven change in the sectoral composition of economies associated with developments in per capita income can give rise to a hump-shaped relationship between per capita income and emissions (see De Groot, 1999). However, and this seems especially relevant for the case of China as discussed previously, the sectoral composition of economies is also affected by patterns of specialisation (between regions or countries). Due to this effect, countries or regions with a comparative advantage in pollution-intensive industries will, *ceteris paribus*, witness an increase in emissions not related to developments in per capita income.⁵

⁴ For China, one could question the relevance of this argument as several Chinese policies are centralised and therefore show limited variation across regions. However, as noted by, e.g., Auffhammer et al. (2001), “some provinces/cities adopted air pollution emission permit policies even before the implementation of any national legislation”.

⁵ Somewhat related to this issue is the recent interest in income inequality as a determining factor. It has been forcefully put forward by, e.g., Ravallion et al. (2000) that distributive issues may play an important role, e.g., because the income elasticity of energy demand is decreasing with income. Since we have no

The hypothesis of the EKC and the ensuing empirical tests have – to say the least – been subject to a fair amount of debate. The empirical tests have been criticised for several reasons. First, there have been criticisms on the way the hypothesis is being tested, using studies for different countries (for example, Dijkgraaf and Vollebergh, 2000, and List and Gallet, 1999). Second, most empirical studies that perform straightforward regression analysis yield relatively little insight into the driving forces that give rise to an EKC. They have at best included time trends to test for developments unrelated to per capita income. These trends may reflect technological progress resulting in lower energy intensities, but they may as well be the resultant of, for example, substitution away from energy in periods of rising energy prices. These problems have been overcome to some extent by decomposition techniques (for example, De Bruyn, 1997, and Sun, 1998). These techniques decompose changes in pollution of energy use into a scale effect, an intensity effect and a structural effect. They thereby give some descriptive idea of the quantitative importance of the factors that may give rise to an EKC.

5. Empirical analysis

Despite the criticisms discussed in the previous section, the empirical assessment of the shape of the Income-Emission Relationship (IER) for China is in our view useful as a descriptive device aimed at detecting some general patterns as well as region-specific effects. In order to allow for the detection of the wide range of potentially relevant functional relationships between income per capita and emissions, we use a flexible specification of the regression equation that allows for linear, quadratic and cubic polynomial relationships between per pollution and per capita income.

5.1 Data

The data set used in our analysis is obtained from official Chinese statistical material (mainly the statistical yearbooks). The data refer to a 16-year time period from 1982–1997, and to the 30 provinces and major cities, to be called regions henceforth. The regions are listed in Table A1 in Appendix 1. For each region, we have land surface and population data. There are figures on total emissions of industrial wastewater, waste gas and solid waste (for every year and region). Moreover, we have data on gross regional product, in current as well as in constant prices.⁶

As a first exploratory analysis of the data, we try to categorise the regions in our data set

data available on income distribution in the regions considered in this study we are unable to explore this matter any further.

⁶ We are aware of the fact that there might be differences between regions with respect to purchasing power. Unfortunately data to correct for this are lacking.

according to their state of economic development and environmental quality for the year 1997. For this aim, we distinguish between regions with low, medium and high pollution levels. We consider wastewater, waste gas and solid waste per square km for each region and express them relative to the averages of the country (larger is 1, smaller is 0). If the sum of the indices exceeds 2 the region is called *highly polluted*, if it is below unity the region is characterised by *low pollution*. Otherwise we call pollution *medium*. We also rank the regions according to income. If one region's GRP per capita exceeds the national average level (6079 yuan), we define the region as an *advanced* region. Otherwise, the region is defined as *underdeveloped*. This results in the classification of regions presented in Table 4 (note that asterisks denote coastal areas, and *N* and *S* indicate North and South, respectively).

< Insert Table 4 about here >

As the table shows, most of the advanced regions are regions with serious pollution, and most of backward regions are medium or light pollution regions. The seriously polluted advanced regions are coastal, and the lightly polluted underdeveloped regions are all landlocked. At first sight, there are no major systematic differences according to the distinction between North and South.

5.2 *The econometric modelling*

Since there exists no consensus about the theoretical background of the Environmental Kuznets Curve, we have considered several specifications of the econometric equation to be estimated. In one type of specification we have total emissions as the dependent variable. This specification is relevant in case the Kuznets curve is driven by consumer preferences that are affected by total pollution rather than by pollution per capita or per unit of production. For some pollutants, such as solid waste, per capita pollution might be relevant. Alternatively, if the engine behind the Kuznets curve is deemed to be production driven, it is more likely that the dependent variable should be related to emissions per unit of production.

For expositional purposes, we distinguish between two basic classes of models that can be estimated. These models all contain the Gross Regional Product per capita (*GRP_CAP*) as the explanatory variable, which can be seen as the key variable of interest in an empirical investigation on the Environmental Kuznets Curve. Extensions to more explanatory variables are straightforward and will be mentioned in the text when they are introduced. The index *i* will denote the region and *t* refers to time.

The first model reads

$$ES_{it} = \alpha_{0i} + \beta_1(GRP_CAP_{it}) + \beta_2(GRP_CAP_{it})^2 + \beta_3(GRP_CAP_{it})^3 + \varepsilon_{it} \quad (1)$$

where ES refers to the pollution indicator. In this model the intercepts are region specific but the income coefficients are uniform.⁷ This model is thus based on the idea that regions follow a similar pattern of development of emissions as they develop, albeit at potentially different levels. These differences in levels can reflect all kinds of unobserved heterogeneity, among which differences in sectoral specialisation are likely to feature prominently.

The second model is more restricted, as it also requires a uniform intercept:

$$ES_{it} = \alpha_0 + \beta_1(GRP_CAP_{it}) + \beta_2(GRP_CAP_{it})^2 + \beta_3(GRP_CAP_{it})^3 + \varepsilon_{it} \quad (2)$$

A general result from the analysis is that the null hypothesis of *equal intercepts* is rejected (using the F-test). This holds for all pollutants and for all specifications thereof (levels, per capita and per unit GRP). Therefore, the second model is not further explored in this paper. All regressions that we present have been estimated with a full set of fixed effects to control for region-specific heterogeneity.⁸

In estimating the models, special care is required in controlling for autocorrelation and heteroscedasticity. Autocorrelation has been addressed by estimating an AR(1) model. To account for heteroscedasticity, we have estimated and reported White's heteroscedasticity consistent estimators. We refer to Greene (1997) for econometric details.

⁷ In principle, one could also allow for different income coefficients and subsequently test for equality (see Dijkgraaf and Vollebergh, 2000). In the present paper we do not pursue this road for two reasons. First, in most EKC studies equality of income coefficients is implicitly assumed. It is the very idea that a genuine relationship exists that forms the motivation to study the shape of the Income Emission Relationship in the first place. The second reason is more fundamental. First, suppose that a "genuine" inverted U shaped EKC exists, uniform with respect to the coefficients corresponding to per capita income, but possibly allowing for heterogeneity in intercepts. Suppose also that a distinction can be made in two groups of regions, one with persistently low per capita income and one with persistently high per capita income. To make the picture complete, assume that there is no overlap in incomes over the observation period. Then tests on equality of income coefficients are likely to reject the null hypothesis (depending on the difference in income between the two sub-groups, the number of observations, the variance of the dependent and the independent variable, etc.). More specifically, the validity of tests on the equality of coefficients declines if the difference in the range of income of different subgroups increases. Therefore, the tests usually employed to test for the homogeneity of coefficients do not enable the researcher to draw strong conclusions on homogeneity.

⁸ We have also tested whether the distinction between North-South and Landlocked-Coastal is statistically sufficient to capture most of the variation in emissions between regions. Except for some special cases, this hypothesis is not supported by the data, so a full set of region-specific fixed effects has been included.

5.3. Wastewater

For water there is a strong negative relationship between per capita income and wastewater emission, measured in levels. The quadratic term as well as the third order variable, although statistically significant, are dominated by the linear term, as is clear from Table 5. There is strong evidence for autocorrelation. A closer inspection of regions with high and low per capita incomes provides some useful information. It turns out that the picture for the five low-income regions is rather fuzzy. The income range for the 18 years under consideration for these regions (being the poorest in 1997) is rather small. Within this small range we observe patterns of both increasing and decreasing wastewater levels over time. For the top 5 richest regions the income range is much larger and the picture shows a clearly decreasing trend over time.

< Insert Table 5 and Figure 1 about here >

Figure 1 graphically depicts the results from the estimation procedure. In the figure, we depict the estimated relationship for the ‘median region’ (that is, the region with the median fixed effect), as well as the relationship for the region with the highest and the lowest fixed effect. The result of a strong negative relationship between emissions and per capita income also obtains when emissions in per capita terms and per unit of GRP are the dependent variables: the linear part is significantly negative and dominates the other terms.

The existence of an environmental Kuznets curve for wastewater is therefore not supported by the data. Instead, a monotonically declining relationship between per capita income and wastewater (in all three dimensions) is found over the relevant range of income levels. Given that the dependent variable is industrial wastewater, a plausible explanation for this result would be that when incomes rise, there is increasing pressure on industries to emit less wastewater. Alternatively, given the strong autocorrelation in all three specifications, and given the very bad initial situation, harassing human health, there might be an autonomous trend in improving water quality over time. Finally, the negative relationship with income could be explained by the fact that the advanced high-income regions specialise relatively strongly in light industries, such as electronics, which are only moderately polluting and at the same time generate higher incomes.

< Insert Figure 2 about here >

Figure 2 contains information about the fixed effects obtained from the regression estimating per capita wastewater emissions. The thirty regions are ranked according to their fixed effects. The

five regions ranking highest are depicted in black, the next five regions in dark-grey, and so on. The fixed effects are largest in the major cities and industrial areas, e.g., Beijing, Shanghai, Tianjin and Jiangsu.

5.4 Waste gas

In the regression with waste gas emissions in levels, all coefficients are statistically significant as shown in Table 6. There is an overall positive relationship between pollution of this type and per capita income, although emissions increase at relatively low rates at intermediate levels of per capita income.

< Insert Table 6 and Figure 3 about here >

When we consider waste gas emissions per capita, the linear income term is statistically significant and dominates: waste gas emissions per capita increase as a function of per capita income. A possible explanation would be that increases in income per capita are strongly related to energy demand, which is in turn strongly related with waste gas emissions. When considering waste gas emissions per unit of output, it turns out again that only the linear term is statistically significant and dominates the other terms. But now the coefficient is negative, indicating that energy intensity of production is decreasing with rising per capita income. Although this change in qualitative relationship is not very large, it is interesting. The fact that emissions of waste gas in levels and per capita increase with per capita income indicates that higher per capita income leads to more energy demand, but that this demand is increasingly associated with energy-intensive products. This is an illustration of the relevance of the composition effect that we discussed in Section 4.

Figure 4 depicts the estimated fixed effects obtained from the regression using waste gas per capita as the dependent variable.

< Insert Figure 4 about here >

As was to be expected, the fixed effects are largest in the North, with its concentration of heavy industries. In 1997, the high-ranking regions (in terms of fixed effects estimated in the equation with emissions in levels) of Guangdong, Shandong, Jiangsu, Hebei and Liaoning jointly produce 34.3% of total China's electricity. Furthermore, they produce 40.3% of China's total output in that year. Waste gas from electricity and cement constituted 36.6% and 13.5% of China's total

industrial waste gas emissions.

5.5 Solid waste

The regression results for industrial solid waste are presented in Table 7.

< Insert Table 7 and Figure 5 about here >

From Table 7 we infer that the coefficients for the equation in levels are statistically highly significant. Moreover, the variation in pollution is well explained. From Figure 5 it is clear that the Income-Emission Relationship is increasing at low per capita income levels as well as at relatively high incomes. It is slightly decreasing for intermediate incomes. Therefore we find support for an N-shaped IER as far as solid waste in levels is concerned. The results in terms of solid waste emissions per capita are completely different. The income coefficients are all statistically insignificant, and there is strong evidence of autocorrelation. To explain this result, one should bear in mind that the dependent variable is not residential solid waste but industrial solid waste, which is more related to income through the scale and intensity effect than through environmental awareness triggering less polluting behaviour. The lack of a clear statistical relationship therefore need not be surprising.

Finally, we have considered solid waste emissions *per unit of GRP* as dependent variable. The coefficients are all statistically significant. However, no inverted U-shaped relation between solid waste per unit of GRP and per capita income can be observed. It appears that the relationship over the relevant domain is negative, suggesting that solid waste emissions per unit of production decrease as incomes increase. Regarding the latter conclusion, it is unclear whether there is a causal relationship or not. Possibly an autonomous mechanism is at work reducing solid waste per unit of production. It is difficult to distinguish between the two possibilities suggested above because income shows a clear time trend.

<Insert Figure 6 about here >

As depicted in Figure 6, the fixed effects obtained from the per capita regression equation differ greatly between regions. They are clearly largest in the northern part of China. It would be constructive to make an effort in explaining the differences. However, given the lack of data no formal econometric analysis can be performed to do so. We confine ourselves to a few observations. For the case of solid waste the province of Liaoning is ranking highest (for all regression specifications we

performed). This can be explained by pointing at the fact that this is a province specialising in the production of raw materials and mining. Iron, steel production and manufacturing of steel products amount to about 11.8%, 12.4% and 11.3% of total production in China. On the other hand, Tianjin has a relatively large secondary industry and has little mining and therefore produces relatively little solid waste (although it is characterised by severe environmental problems on aggregate, as shown in the remainder of this section). Shanxi produces a considerable amount of coal (nearly 1/4 of China's output in 1997), so it produced more solid waste.

5.6 Prediction

It is tempting to predict future Chinese emissions of pollutants based on the previous econometric analysis. In particular, given the intensive discussion regarding the enhanced greenhouse effect and the relevance of participation of less developed countries, predicting the emissions of greenhouse gases would be of interest. Such exercises have been performed before (e.g., Martins et al., 1992, Manne and Richels, 1992, and Auffhammer et al., 2001). Manne and Richels (1992) estimate China's annual growth rate of CO₂ emissions over 1990-2050 at about 2.3%. Martins et al. (1992) estimate a growth rate of 3.7%. The former suggests that China's share in CO₂ in the world will rise from 9.5% in 1985 to 17% in 2050; the latter estimates the share to be 29% in 2050. Auffhammer et al. (2001) also report on CO₂ emissions, but now in total volume in 2050, and compare the results with other predictions, based on an IER approach.

Since we have data on total industrial waste gas only, it is difficult to compare our results with those mentioned above. Moreover, as a general remark, care is required in extrapolating the regression results far into the future as there is no guarantee that the estimated relationships are stable over time, nor that the estimation results are representative for income levels that far exceed those that are represented in the sample. Nevertheless, we have made some tentative projections for waste gas emissions, based on our regional analysis of the IER. We have extrapolated our regression results for waste gas into the future assuming different scenarios for the growth rate of GRP per capita. The scenarios and their implications for emissions are given in Table 8.

< Insert Table 8 about here >

In all scenarios we assume that the growth rate of GRP per capita linearly converges to 2% in 2050. Our scenarios account for different initial growth rates and for potentially heterogeneous development of the coastal and landlocked regions.

In the pessimistic development scenario, growth rates of GRP per capita equal 2% per annum throughout the first half of the 21st century. They are uniform for the coastal and the landlocked regions. The predicted growth rate of emissions then slightly exceeds the growth rate of GRP per capita and is predicted to equal 2.36%. In the optimistic scenario, both the coastal and landlocked regions start at an annual growth rate of 5% that converges to 2% in 2050. Emissions are then predicted to grow at 7.30% and thus far exceed the growth of GRP per capita. This mainly reflects the N-shaped Income Emission Relationship that predicts emissions to grow at an accelerating rate as countries pass a critical level of GRP per capita. Finally, we consider two scenarios of unequal development of the coastal and landlocked regions. In the divergence scenario, the coastal region develops as in the optimistic development scenario, while the landlocked regions only grow at 3% in 2000. This results in a minor reduction in the growth of emissions to 7.27% reflecting the fact that the landlocked regions start at relatively low levels of GRP per capita and by 2050 have not yet surpassed the critical level of GRP per capita above which emissions start to increase drastically in any of the two development scenarios. In contrast, the convergence scenario in which the landlocked regions develop according to the optimistic scenario, whereas the coastal regions start at 3% GRP growth in 2000 results in a sharp decline of the growth rate of emissions to 3.97%, for opposite reasons as those discussed in the divergence scenario.

Two remarks on these results are in place when we compare them to other predictions. First, given the sharp decline in population growth in China that has taken place recently and that is projected to continue for the next few decades, and given the fact that we did not include population size as a separate explanatory variable in the regression results that we have reported in this section, our predictions are likely to be on the high side of the relevant range.⁹ Second, our regression results have revealed the relevance of N-shaped Income Emission Relationships. Such relationships result in accelerating emissions after regions have surpassed a certain per capita income level. By construction, such relationships are absent in Aufhammer et al. (2001) since they only include a squared per capita income term. Their predictions therefore tend to underestimate the growth rate of emissions - almost by construction - if the N-shaped IER is indeed the empirically relevant one as our results suggest.

Clearly, given the importance of China's future CO₂ emissions in the context of the enhanced

⁹ We have tested for the effects of including population as an explanatory variable, but this did not substantially affect the results reported in this section on the IER. Insofar the development of GDP and population are correlated, the omission of population is also not a problem for the quality of the prediction. If, however, the relationship will change, the quality of the predictions will fall because of parameter instability over time.

greenhouse effect, much more study is required before strong conclusions can be drawn that can be used in the debate on global CO₂-emissions. Still, we are convinced that the approach adopted in this paper that explicitly distinguishes Chinese regions and their potentially heterogeneous development is relevant for improving the quality of predictions (see also Auffhammer et al., 2001). In any case, all results that have been obtained so far point at drastic increases of China's emissions as a fraction of total world emissions and therefore indicate the relevance of China and its future development for understanding, predicting and ultimately curing the Chinese future contribution to the enhanced greenhouse gas effect.

6. Conclusions and policy recommendations

In this paper, we have analysed China's regional development and the associated change of China's regional environment. The main finding is that clear relationships between Gross Regional Product and pollution exist. The shape of the relationship heavily depends on the pollutant that is considered and on the specification of the dependent variable (in levels, in per capita terms or per unit of GRP). For wastewater, we find a monotonic negative relationship indicating that water quality increases with income (and time). This holds for all specifications. For waste gas in levels there is a typical Kuznets pattern. When waste gas emissions are modelled in per capita terms, the relationship is monotonically increasing in per capita income, although at a rate that depends on per capita income. When they are measured per unit of output, the relationship is decreasing. For solid waste we find support for an N-shaped Income Emission Relationship when estimated in levels. For solid waste per capita, no statistically significant relationship with per capita income is found.

The fixed effects describing differences between regions are statistically highly significant and imply that there are major differences between regions. These differences are likely to reflect differences in the underlying industrial structure. More in-depth statistical analysis would be required to further substantiate this conclusion. Lack of suitable data does not allow us to perform such an analysis at this stage.

We have also made some tentative predictions of waste gas emissions, based on the regional IER analysis. It turns out that even for exogenously determined moderate per capita growth rates the increase in emissions is considerable at 7% annually on average until 2050. The fact that this increase is higher than the per capita growth rate follows from catching up of the poorer regions and from the N-shaped Income Emission Relationship that suggests accelerating emissions after a certain per capita income level is surpassed. Fundamental transformation in production processes

and exploitation of technological opportunities that are available in the West reducing the energy intensity of production are therefore of utmost importance.

In designing environmental policies, China will have to take account of its own special position. But it might also benefit from experiences with environmental policies elsewhere in the world. There might be some room for economic instruments such as levies and permits, which seem to be suitable in an economy where market forces are already in place and will play a more dominant role in the future. The shortcomings of government intervention and market failure contribute to China's pollution problem. This occurred inevitably in China's economic transformation with loosening economic planning and intervention. The important question to care about now is not to distinguish the beneficiaries from the sufferers of this environmental degradation. It is important to determine who is responsible for pollution damage, who is responsible for environmental improvement, and how to join forces to curb pollution by institutional management so as to get high benefits at low cost. In these aspects, China could learn lessons from advanced countries, and also from its own experiences. We end by briefly describing some threats and opportunities for Chinese environmental policy.

In general, China's pollution situation is critical, especially in advanced regions. Some types of pollution are still increasing. Although the central government recognises the imperative of environmental protection, Jiang Zemin has put forward that in the process of economic development "we can't eat the food the ancestors inherited, and cut the road of descendants for living". But with an enormous population, relative scarcity of natural resources, and an ambitious development plan, it is "more difficult to have effective environmental protection" (Jan, 1995). Therefore the outlook is not positive. But it should be noticed that there are factors constraining environmental deterioration. The commodity market turns its weight from seller-preference market to buyer-preference market; the reduction of growth rates diminishes demand for energy consumption; at present there is excess supply of electricity and coal. Moreover, China has drawn up dozens of laws and acts linked with environmental protection, issued over 300 environmental standards. Local governments even promulgated over 600 rules to protect the environment. There are over 100 thousand people involved in environmental management and environmental studies (National Environment Protection Agency, 1996). The central government controls the pollution spread of Huai River and Tai Lake successfully. More and more people are concerned about their production and living environment. Only if China pays additional attention to environmental protection, it can offset the pressure from population, scarce resources and development.

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Table 1. Share (in %) of GDP and GDP per capita of coastal areas in total

| Year | GDP | GDP per capita |
|------|------|----------------|
| 1978 | 52.6 | 61.6 |
| 1980 | 52.3 | 61.4 |
| 1986 | 53.1 | 61.7 |
| 1991 | 55.1 | 63.5 |
| 1995 | 58.3 | 66.6 |
| 1997 | 58.0 | 66.3 |

Source: Processed from China's Regional Historic Statistical Material Collection (1949-1989) and China Statistical Yearbooks (1992), (1996) and (1998).

Table 2. Northern share in some economic indicators

| | 1978 | 1985 | 1990 | 1996 |
|----------------------------------|------|------|------|------|
| Population | 42.1 | 42.2 | 42.3 | 42.2 |
| Area | 59.3 | 59.3 | 59.3 | 59.3 |
| GDP | 46.5 | 44.6 | 43.9 | 41.1 |
| Grain output | 40.9 | 40.7 | 44.6 | 48.2 |
| Gross output value of industry | 49.4 | 45.1 | 44.5 | 39.5 |
| Total investment in fixed assets | 51.5 | 50.0 | 46.6 | 38.4 |

Source: Data processed from China's Regional Historic Statistical Material Collection (1949-1989) and China Statistical Yearbooks.

Table 3. Industrial pollution in coastal areas (absolute and relative to total industrial pollution)

| Pollution type | Year | Coastal areas absolute | Coastal areas % of total |
|-------------------------------|------|---------------------------|-----------------------------|
| Waste water (million tons) | 1982 | 11674.3 | 50.0% |
| | 1985 | 11717.7 | 47.1% |
| | 1990 | 12150.4 | 47.0% |
| | 1995 | 11145.5 | 50.2% |
| | 1997 | 9390.4 | 49.7% |
| Waste gas (billion m^3) | 1982 | 2264.7 | 47.2% |
| | 1985 | 2745.8 | 45.5% |
| | 1990 | 3596.7 | 49.2% |
| | 1995 | 5242.1 | 49.7% |
| | 1997 | 5734.1 | 50.6% |
| Solid waste (million tons) | 1982 | 173.7 | 43.4% |
| | 1985 | 208.6 | 42.9% |
| | 1990 | 253.2 | 43.8% |
| | 1995 | 282.7 | 43.7% |
| | 1997 | 279.0 | 42.5% |

Source: China's Environmental Statistical Materials (1981-90) and China Statistical Yearbook (1991, 1996, 1998).

Table 4. Regional ordering based on pollution and economic development

| | <i>Serious pollution</i> | <i>Medium pollution</i> | <i>Light pollution</i> |
|-------------------------|--|--|---|
| <i>Advanced regions</i> | Beijing*(N) Liaoning*(N) Tianjin*(N) Shandong*(N) Guangdong*(S) Jiangsu*(S) Shanghai*(S) Zhejiang*(S) | Fujian*(S) Heilongjiang(N) | |
| <i>Backward regions</i> | Hebei*(N) Henan(N) Shanxi(N) Anhui(S) | Guangxi*(S), Hainan*(S) Jilin(N) Ningxia(N) Shaanxi(N) Guizhou(S) Hubei(S) Hunan(S) Jiangxi(S) Sichuan(S) | Gansu(N) Neimenggu(N) Qinghai(N) Xinjiang(N) Yunnan(S) Xizang(S) |

Table 5. Parameter estimates for wastewater (sample period 1982-1997)

| Dependent variable | Wastewater in levels | | Wastewater per capita | | Wastewater per unit of GRP | |
|------------------------|------------------------|-------------------|--------------------------|-------------------|----------------------------|-------------------|
| | Estimate | White t-statistic | Estimate | White t-statistic | Estimate | White t-statistic |
| GRP_CAP | -9.99E+00 [†] | -5.75 | -2.00E-03 ^{***} | -3.64 | -2.33E-02 ^{**} | -7.25 |
| (GRP_CAP) ² | 5.67E-04 [*] | 3.74 | 7.50E-08 | 1.22 | 1.40E-06 ^{***} | 5.15 |
| (GRP_CAP) ³ | -1.28E-08 [*] | -3.31 | -3.24E-12 ^{**} | -1.97 | -2.88E-11 ^{***} | -4.40 |
| AR(1) | 0.65 ^{***} | 10.71 | 0.71 ^{***} | 16.79 | 0.85 ^{***} | 34.06 |
| Beijing | 96237 | | 53.7 | | 150.2 | |
| Tianjin | 74042 | | 40.1 | | 140.7 | |
| Hebei | 118779 | | 21.2 | | 100.4 | |
| Shanxi | 73751 | | 22.1 | | 105.4 | |
| Neimenggu | 53037 | | 18.1 | | 100.4 | |
| Liaoning | 188063 | | 46.8 | | 139.9 | |
| Jilin | 82069 | | 27.4 | | 103.3 | |
| Heilongjiang | 122816 | | 32.9 | | 135.6 | |
| Shanghai | 197471 | | 125.2 | | 195.8 | |
| Jiangsu | 257578 | | 42.0 | | 143.5 | |
| Zhejiang | 140339 | | 34.1 | | 135.0 | |
| Anhui | 115049 | | 21.3 | | 110.7 | |
| Fujian | 102460 | | 30.9 | | 128.9 | |
| Jiangxi | 92064 | | 22.9 | | 106.3 | |
| Shandong | 129060 | | 19.6 | | 111.2 | |
| Henan | 128640 | | 17.5 | | 92.5 | |
| Hubei | 174739 | | 33.1 | | 124.2 | |
| Hunan | 193350 | | 32.9 | | 135.4 | |
| Guangdong | 171433 | | 30.7 | | 120.0 | |
| Guangxi | 114396 | | 25.7 | | 107.7 | |
| Hainan | 40873 | | 18.6 | | 101.6 | |
| Sichuan | 218215 | | 22.9 | | 110.4 | |
| Guizhou | 49041 | | 13.2 | | 66.6 | |
| Yunnan | 66589 | | 16.6 | | 96.9 | |
| Xizang | 22358 | | 11.4 | | 90.9 | |
| Shaanxi | 60214 | | 16.1 | | 85.0 | |
| Gansu | 55478 | | 19.8 | | 88.9 | |
| Qinghai | 31774 | | 18.5 | | 89.1 | |
| Ningxia | 32807 | | 23.5 | | 114.4 | |
| Xinjiang | 48712 | | 18.2 | | 105.5 | |
| R ² | 0.98 | | 0.99 | | 0.97 | |
| Observations | 442 | | 442 | | 442 | |
| F-statistic | 7891.2 ^{***} | | 10930.6 ^{***} | | 3959.0 ^{***} | |
| F-statistic | 3.13 ^{***} | | 2.03 ^{***} | | 1.12 | |
| fixed effects | | | | | | |

Note: The t-statistics that are reported are White-heteroskedasticity consistent t-statistics. They are robust to heteroskedasticity within each cross-section, but do not account for the possibility of contemporaneous correlation across cross-sections. Significance levels are indicated with stars: ***, **, and * means significance at 1, 5 and 10%, respectively. The fixed effects are reported and have been tested for joint significance. The F-statistic fixed effects tests for the joint significance of the fixed effects (with the null-hypothesis that all fixed effects are equal to zero). All models estimated without fixed effects are significant at 1% significance level.

Table 6. Parameter estimates for industrial waste gas emissions (sample 1982-1997)

| Dependent variable | Gas in levels | | Gas per capita | | Gas per unit of GRP | |
|------------------------|--------------------------|-------------------|-------------------------|-------------------|--------------------------|-------------------|
| | Estimate | White t-statistic | Estimate | White t-statistic | Estimate | White t-statistic |
| GRP_CAP | 8.70E-01 ^{***} | 7.95 | 1.47E-04 ^{***} | 5.48 | -4.62E-04 ^{***} | -4.11 |
| (GRP_CAP) ² | -5.04E-05 ^{***} | -4.72 | -7.06E-09 ^{**} | -2.28 | 2.48E-08 ^{**} | 2.28 |
| (GRP_CAP) ³ | 1.00E-09 ^{***} | 3.46 | 1.65E-13 [*] | 1.77 | -5.09E-13 [*] | -1.79 |
| AR(1) | 0.51 ^{***} | 2.97 | 0.13 | 0.83 | 0.11 | 0.68 |
| Beijing | -2005 | | 1.34 | | 5.40 | |
| Tianjin | -2709 | | 0.79 | | 4.61 | |
| Hebei | 3211 | | 0.47 | | 4.07 | |
| Shanxi | 1767 | | 0.93 | | 5.77 | |
| Neimenggu | 477 | | 0.77 | | 5.02 | |
| Liaoning | 4138 | | 1.28 | | 5.68 | |
| Jilin | 56 | | 0.57 | | 4.37 | |
| Heilongjiang | 689 | | 0.53 | | 4.07 | |
| Shanghai | -1361 | | 1.54 | | 5.31 | |
| Jiangsu | 2032 | | 0.20 | | 3.35 | |
| Zhejiang | -628 | | 0.02 | | 2.94 | |
| Anhui | 851 | | 0.15 | | 3.05 | |
| Fujian | -1532 | | -0.06 | | 2.61 | |
| Jiangxi | -89 | | 0.15 | | 3.06 | |
| Shandong | 2975 | | 0.18 | | 3.18 | |
| Henan | 3024 | | 0.25 | | 3.61 | |
| Hubei | 828 | | 0.16 | | 3.07 | |
| Hunan | 735 | | 0.11 | | 2.88 | |
| Guangdong | 891 | | 0.04 | | 2.98 | |
| Guangxi | 19 | | 0.09 | | 2.56 | |
| Hainan | -2610 | | -0.24 | | 2.02 | |
| Sichuan | 3097 | | 0.12 | | 2.96 | |
| Guizhou | 409 | | 0.28 | | 3.94 | |
| Yunnan | -566 | | 0.03 | | 2.51 | |
| Xizang | -1719 | | -0.25 | | 1.04 | |
| Shaanxi | -136 | | 0.19 | | 3.12 | |
| Gansu | 530 | | 0.67 | | 6.09 | |
| Qinghai | -1667 | | 0.69 | | 5.22 | |
| Ningxia | -1360 | | 0.88 | | 5.57 | |
| Xinjiang | -1445 | | 0.28 | | 3.25 | |
| R ² | 0.92 | | 0.83 | | 0.52 | |
| Observations | 441 | | 441 | | 441 | |
| F-statistic | 1639.1 ^{***} | | 645.9 ^{***} | | 144.5 ^{***} | |
| F-statistic | 4.23 ^{***} | | 7.84 ^{***} | | 5.5 ^{***} | |
| fixed effects | | | | | | |

Note: The t-statistics that are reported are White-heteroskedasticity consistent t-statistics. They are robust to heteroskedasticity within each cross-section, but do not account for the possibility of contemporaneous correlation across cross-sections. Significance levels are indicated with stars: ^{***}, ^{**}, and ^{*} means significance at 1, 5 and 10%, respectively. The fixed effects are reported and have been tested for joint significance. The F-statistic fixed effects tests for the joint significance of the fixed effects (with the null-hypothesis that all fixed effects are equal to zero). All models estimated without fixed effects are significant at 1% significance level.

Table 7. Parameter estimates for industrial solid waste (sample 1982-1997)

| Dependent variable | Solid waste in levels | | Solid waste per capita | | Solid waste per unit of GRP | |
|------------------------|-----------------------|-------------------|------------------------|-------------------|-----------------------------|-------------------|
| | Estimate | White t-statistic | Estimate | White t-statistic | Estimate | White t-statistic |
| GRP_CAP | 1.65E-01*** | 3.36 | 1.08E-05 | 0.91 | -6.07E-04*** | -10.18 |
| (GRP_CAP) ² | -1.10E-05*** | -2.67 | -3.46E-11 | -0.03 | 4.08E-08*** | 7.43 |
| (GRP_CAP) ³ | 2.38E-10*** | 2.40 | -1.68E-15 | -0.05 | -8.56E-13*** | -5.93 |
| AR(1) | 0.73*** | 12.74 | 0.64*** | 12.94 | 0.67*** | 14.40 |
| Beijing | 137 | | 0.68 | | 3.61 | |
| Tianjin | -288 | | 0.40 | | 3.25 | |
| Hebei | 5224 | | 0.86 | | 4.45 | |
| Shanxi | 3367 | | 1.21 | | 5.59 | |
| Neimenggu | 1729 | | 0.91 | | 4.55 | |
| Liaoning | 6569 | | 1.74 | | 5.57 | |
| Jilin | 1059 | | 0.58 | | 3.36 | |
| Heilongjiang | 2941 | | 0.92 | | 4.20 | |
| Shanghai | 328 | | 0.70 | | 3.55 | |
| Jiangsu | 1829 | | 0.29 | | 2.83 | |
| Zhejiang | 331 | | 0.15 | | 2.61 | |
| Anhui | 2266 | | 0.42 | | 3.19 | |
| Fujian | 136 | | 0.17 | | 2.58 | |
| Jiangxi | 3139 | | 0.84 | | 5.03 | |
| Shandong | 3480 | | 0.41 | | 3.02 | |
| Henan | 2030 | | 0.24 | | 2.40 | |
| Hubei | 1476 | | 0.31 | | 2.77 | |
| Hunan | 1510 | | 0.28 | | 2.63 | |
| Guangdong | 948 | | 0.18 | | 2.75 | |
| Guangxi | 1003 | | 0.28 | | 2.59 | |
| Hainan | -435 | | 0.11 | | 2.34 | |
| Sichuan | 3637 | | 0.34 | | 2.82 | |
| Guizhou | 804 | | 0.29 | | 2.78 | |
| Yunnan | 1593 | | 0.49 | | 3.40 | |
| Xizang | -346 | | -0.02 | | 1.28 | |
| Shaanxi | 1347 | | 0.46 | | 3.22 | |
| Gansu | 941 | | 0.50 | | 3.57 | |
| Qinghai | -170 | | 0.50 | | 3.35 | |
| Ningxia | -23 | | 0.74 | | 4.15 | |
| Xinjiang | -9 | | 0.27 | | 2.62 | |
| R ² | 0.98 | | 0.97 | | 0.96 | |
| Observations | 442 | | 442 | | 442 | |
| F-statistic | 7648.4*** | | 4693.2*** | | 2913.6*** | |
| F-statistic | 4.12*** | | 4.74*** | | 3.04*** | |
| fixed effects | | | | | | |

Note: The t-statistics that are reported are White-heteroskedasticity consistent t-statistics. They are robust to heteroskedasticity within each cross-section, but do not account for the possibility of contemporaneous correlation across cross-sections. Significance levels are indicated with stars: ***, **, and * means significance at 1, 5 and 10%, respectively. The fixed effects are reported and have been tested for joint significance. The F-statistic fixed effects tests for the joint significance of the fixed effects (with the null-hypothesis that all fixed effects are equal to zero). All models estimated without fixed effects are significant at 1% significance level.

Table 8. Predicted growth of waste gas emissions for different development scenario's

| | Pessimistic | Optimistic | Optimistic - Divergence | Optimistic - Convergence |
|--|-------------|------------|----------------------------|-----------------------------|
| Growth rates in 2000 and 2050 (under assumption of linear convergence) | | | | |
| • Coast | 2% - 2% | 5% - 2% | 5% - 2% | 3% - 2% |
| • Landlocked | 2% - 2% | 5% - 2% | 3% - 2% | 5% - 2% |
| Predicted annual emission growth rate 2000-2050 | | | | |
| | 2.36% | 7.30% | 7.27% | 3.97% |

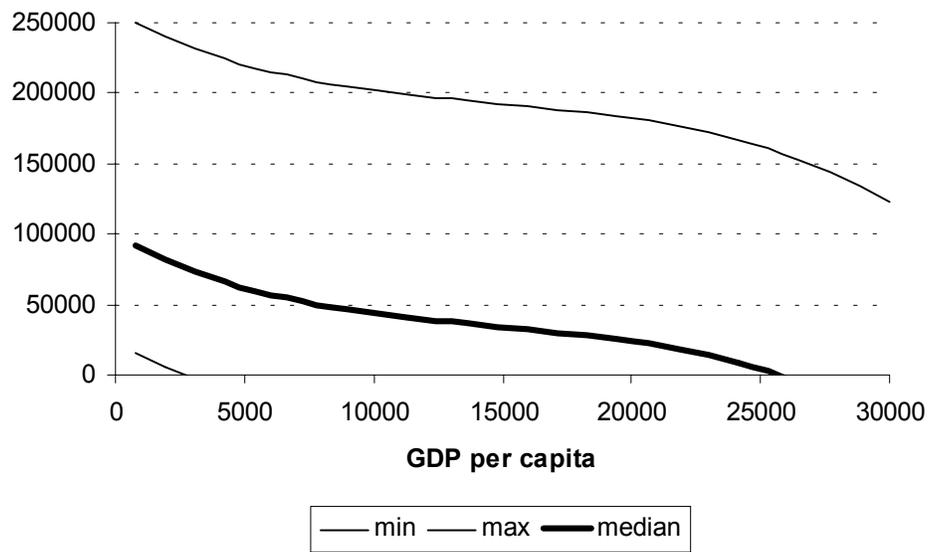


Figure 1. Industrial wastewater as a function of per capita GRP.

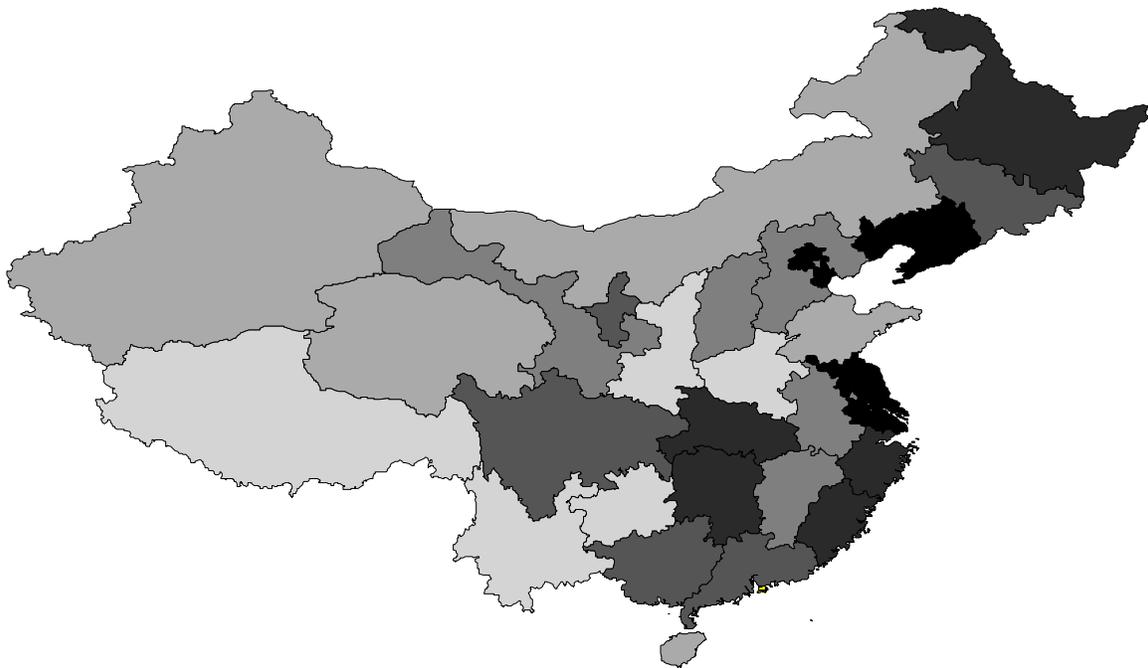


Figure 2. Estimated fixed effects: industrial wastewater per capita

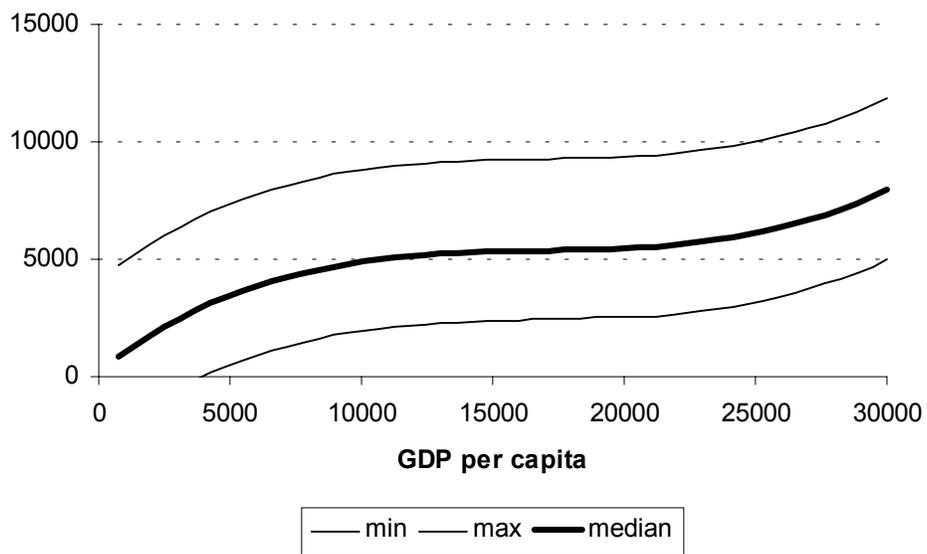


Figure 3. Industrial waste gas emissions as a function of per capita GRP.

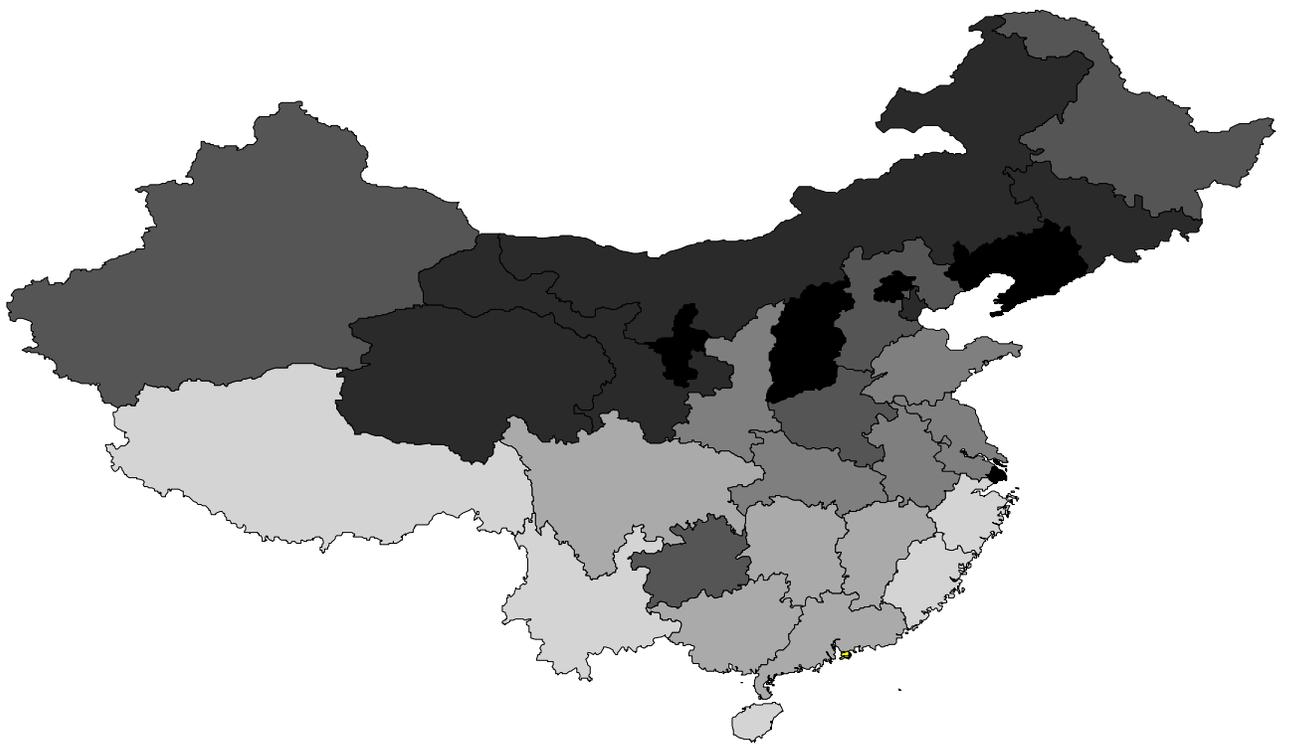


Figure 4. Estimated fixed effects: Industrial waste gas per capita

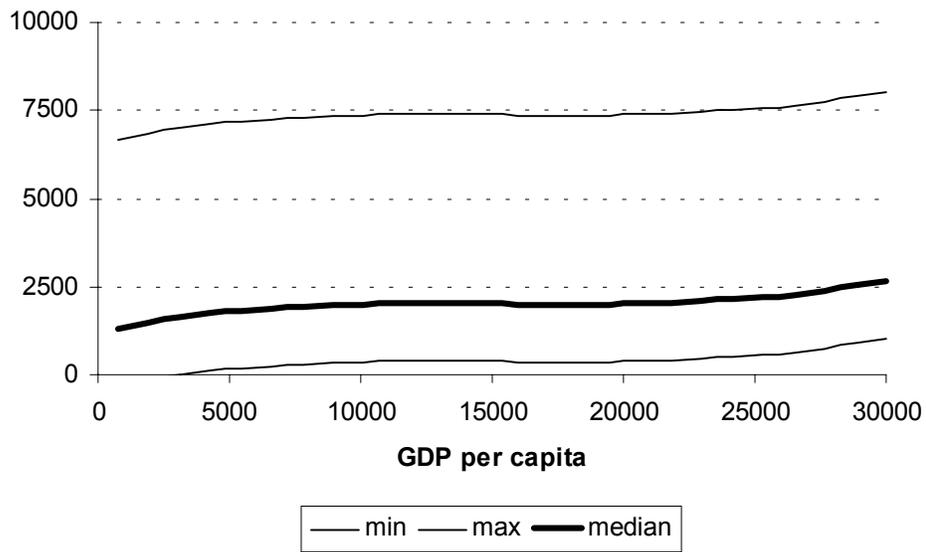


Figure 5. Solid waste as a function of per capita GRP.

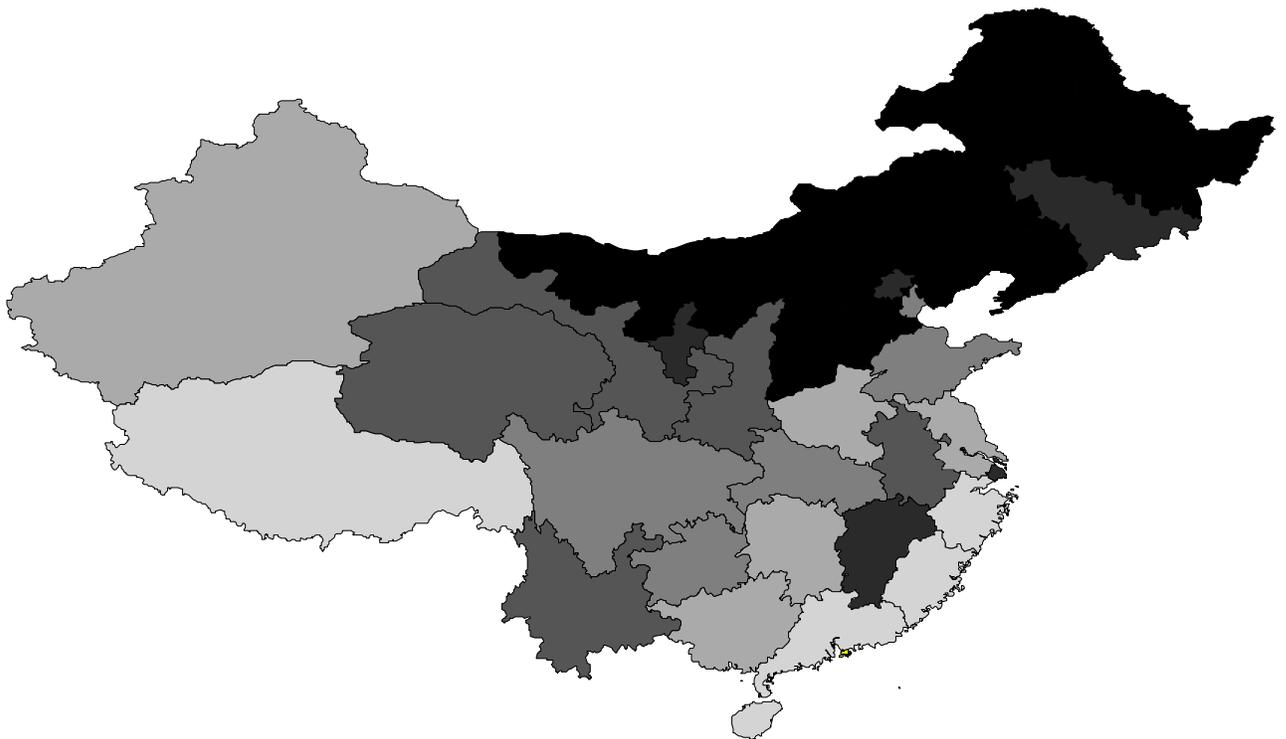


Figure 6. Estimated fixed effects: solid waste per capita

Appendix 1

Table A1. Comparison of average regional growth rates (%) between 1981-92 and 1994-97¹⁰

| Region | Annual growth rate of GRP 1978-1993 | Annual growth of GRP 1994-1997 | Annual growth rate of per capita GRP 1978-97 |
|------------------|-------------------------------------|--------------------------------|--|
| Beijing* (N) | 9.4 | 11.2 | 9.4 |
| Tianjin* (N) | 8.1 | 13.9 | 9.0 |
| Hebei* (N) | 9.7 | 13.7 | 10.2 |
| Liaoning* (N) | 8.7 | 8.9 | 8.6 |
| Shanghai* (S) | 8.4 | 13.5 | 9.5 |
| Jiangsu* (S) | 12.3 | 14.0 | 12.9 |
| Zhejiang* (S) | 13.3 | 15.5 | 14.2 |
| Fujian* (S) | 13.3 | 16.7 | 13.7 |
| Shandong* (N) | 11.5 | 13.5 | 12.0 |
| Guangdong* (S) | 13.9 | 13.8 | 13.3 |
| Guangxi* (S) | 9.2 | 12.6 | 9.0 |
| Hainan* (S) | - | 7.9 | - |
| Shanxi (N) | 8.5 | 10.5 | 6.7 |
| Neimenggu (N) | 9.8 | 10.4 | 9.5 |
| Jilin (N) | 9.2 | 11.2 | 9.6 |
| Heilongjiang (N) | 6.9 | 9.7 | 7.2 |
| Anhui (S) | 9.8 | 15.5 | 10.6 |
| Jiangxi (S) | 9.8 | 14.1 | 10.1 |
| Henan (N) | 10.4 | 13.2 | 10.6 |
| Hubei (S) | 9.9 | 14.0 | 10.4 |
| Hunan (S) | 8.4 | 11.2 | 8.5 |
| Sichuan (S) | 9.2 | 10.4 | 9.9 |
| Guizhou (S) | 9.2 | 8.6 | 8.3 |
| Yunnan (S) | 9.7 | 10.6 | 9.2 |
| Xizang (S) | - | 14.5 | - |
| Shaanxi (N) | 9.6 | 9.3 | 8.9 |
| Gansu (N) | 8.4 | 10.1 | 8.8 |
| Qinghai (N) | 6.6 | 8.5 | 5.9 |
| Ningxia (N) | 8.9 | 10.6 | 7.9 |
| Xinjiang (N) | 11.2 | 9.3 | 9.8 |

Source: Data on GRP (1978-1993) from The Gross Domestic Product of China (1952-1995) and China Statistical Yearbook (1998).

¹⁰ Asterisks denote coastal regions.

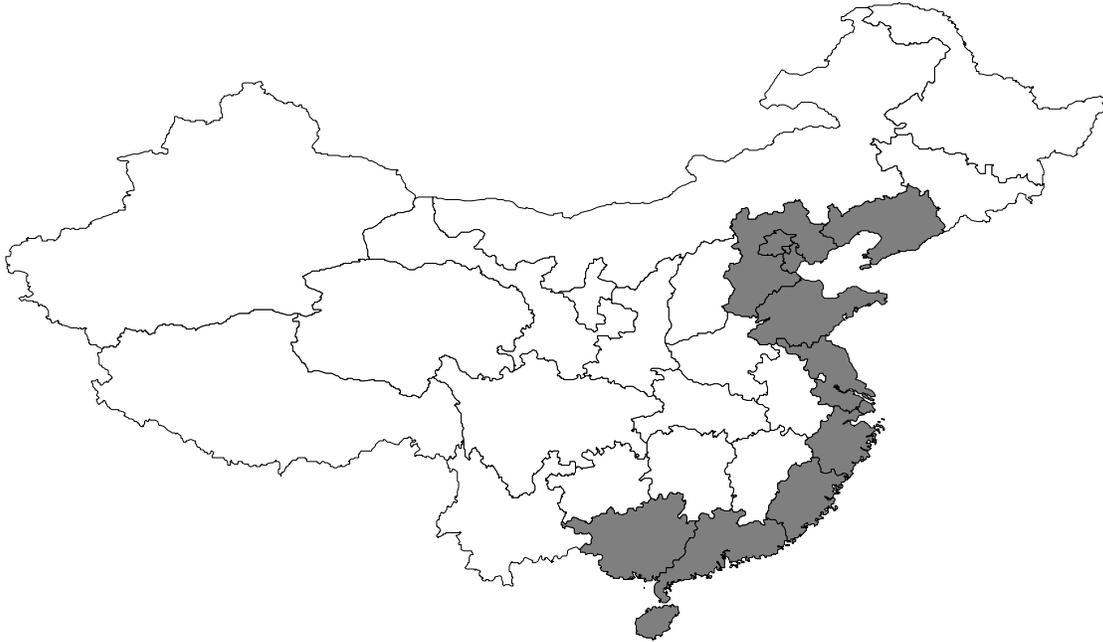


Figure A1 Coast versus non-coast

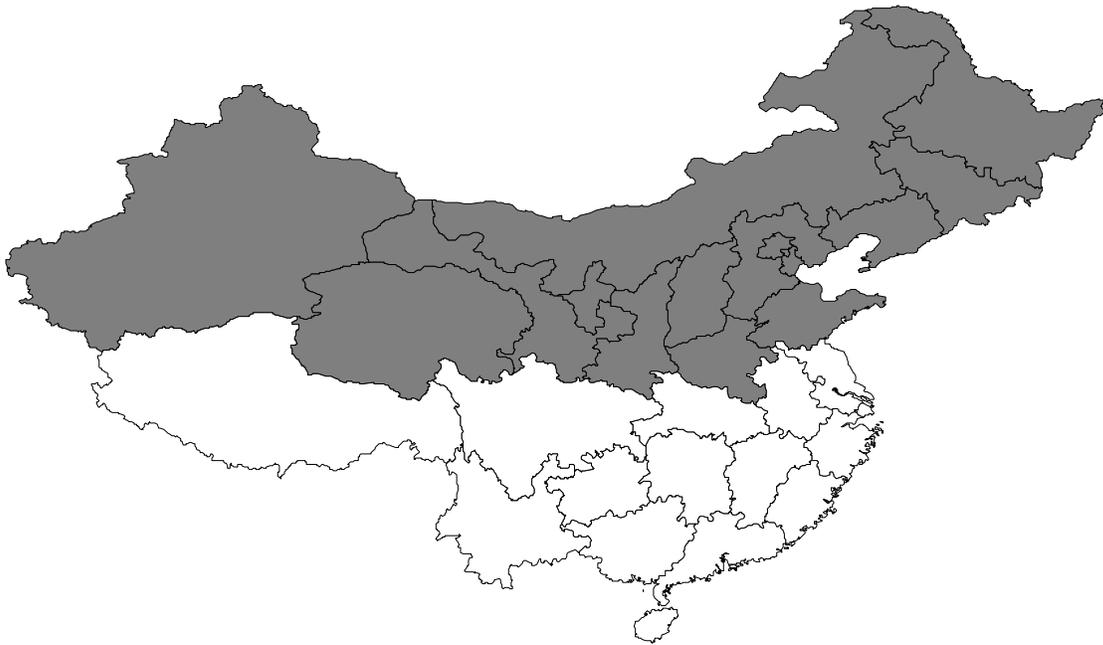


Figure A2 North versus South



Figure A3 Names of provinces

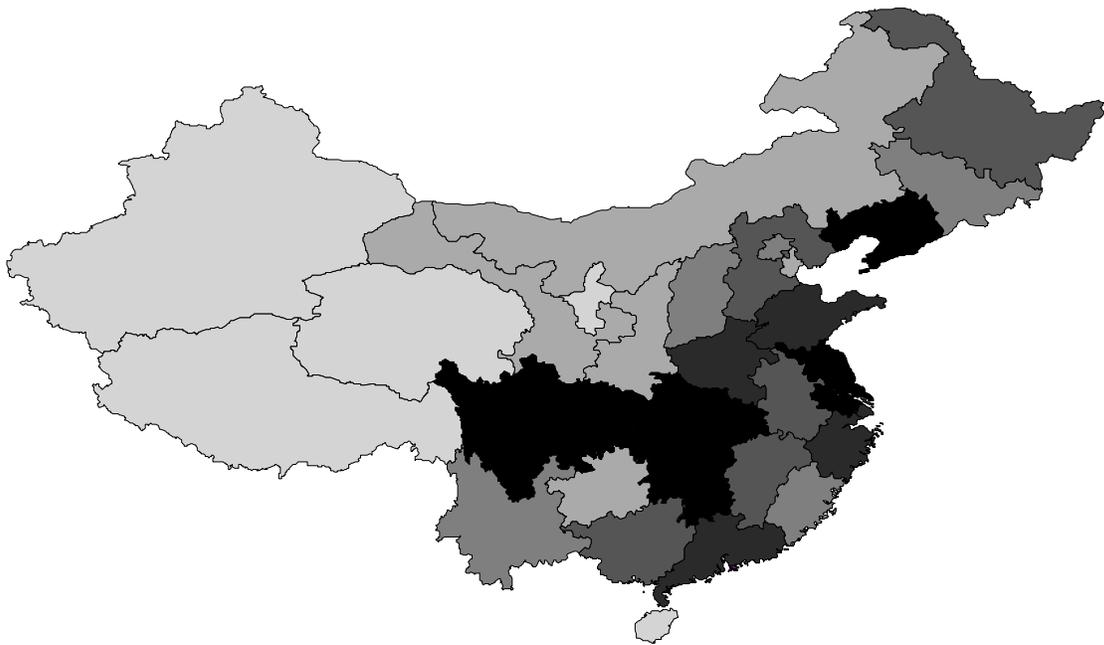


Figure A4 Wastewater emissions in levels in 1997

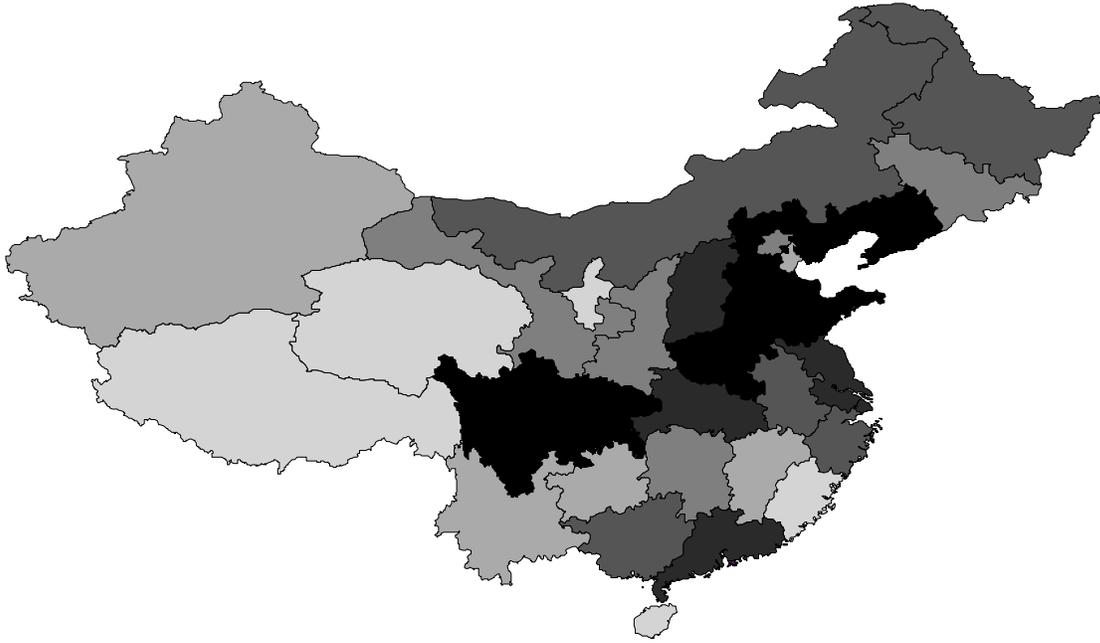


Figure A5 Gas emissions in levels in 1997

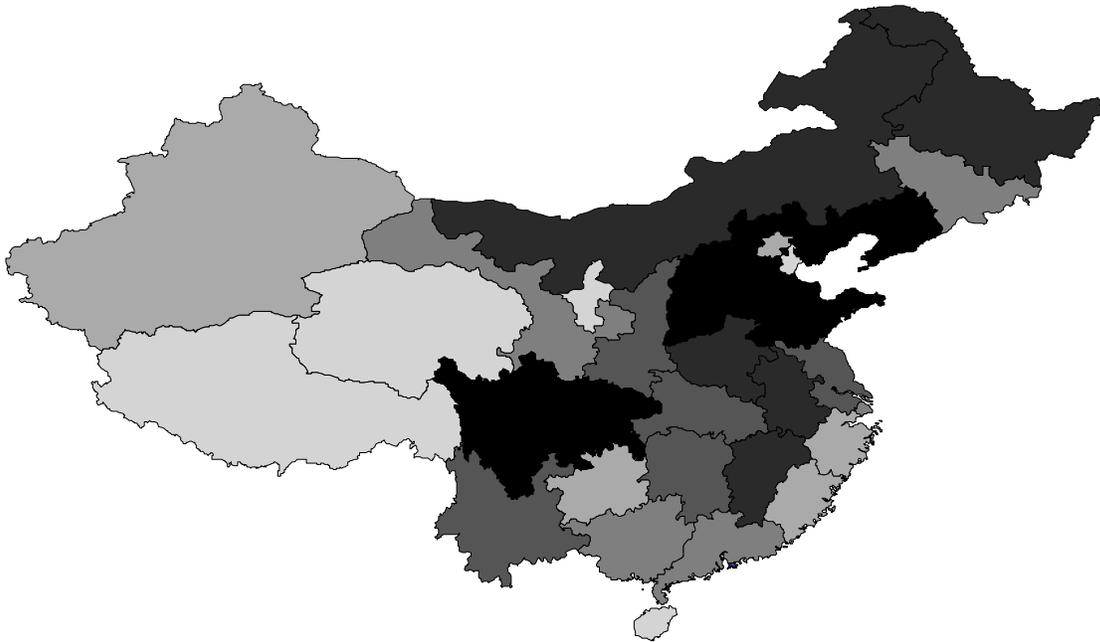


Figure A6 Solid emissions in levels in 1997